

# Laplace approximations in double additive cure survival models with exogenous time-varying covariates

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# Cure survival models

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- In **classical survival analysis**, it is often implicitly assumed that all units will experience the event of interest if one waits for sufficiently long.
- Mathematically, this translates into a population survival function equal to zero at infinity:

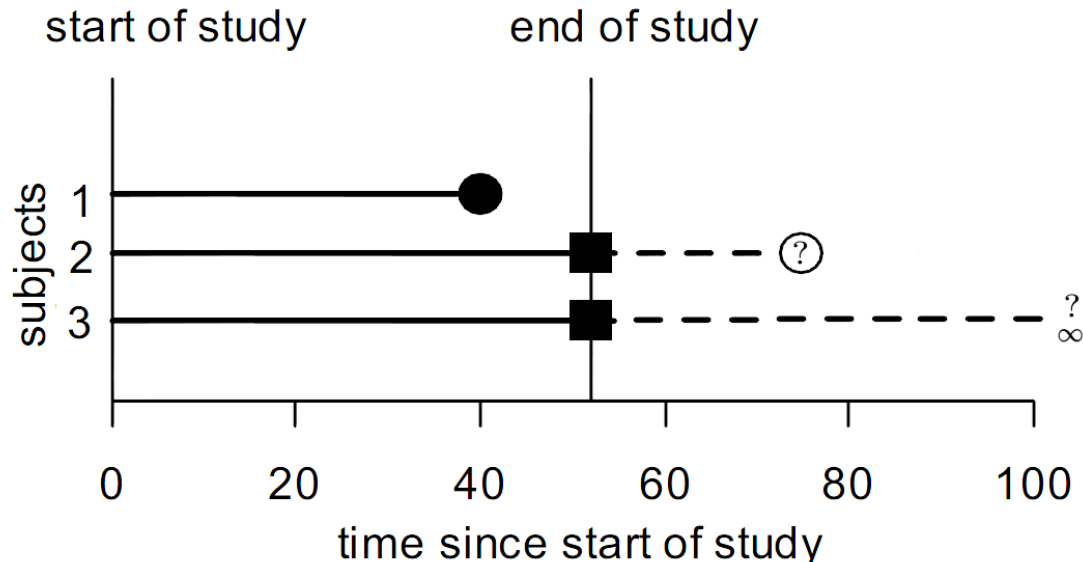
$$S_p(+\infty) = \lim_{t \rightarrow +\infty} \Pr(T > t) = 0$$

- However, in many practical situations, an unidentified **cure fraction** of the population will never experience the event of interest:
  - ▷ Immunotherapy: patients under treatment might be totally cured when the immune system is reactivated.
  - ▷ Fertility study: women under follow-up might have decided not to have a child.
- In these situations,  $S_p(\cdot)$  is an improper (population) survival function:

$$S_p(+\infty) = \pi > 0 \quad (\text{Long-term survival})$$

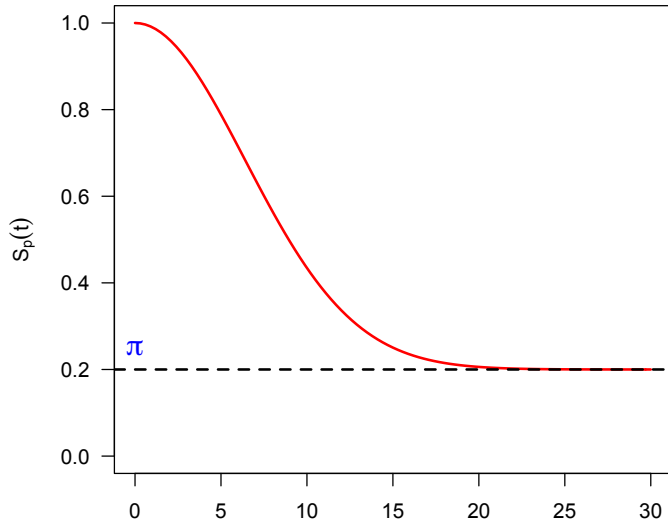
# Right censoring and cure models

- The major difficulty when a cure fraction is present is the status of subjects with a right-censored response.
- Will a subject with a right-censored response finally experience the event ?

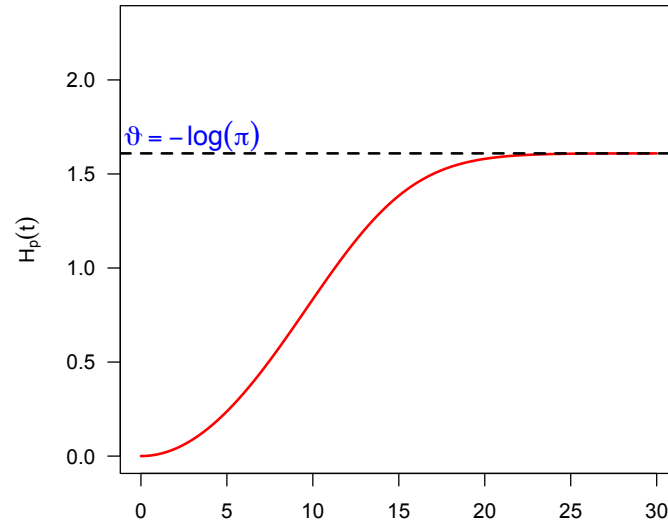


# The promotion time cure model

Population survival function



Population cumulative hazard

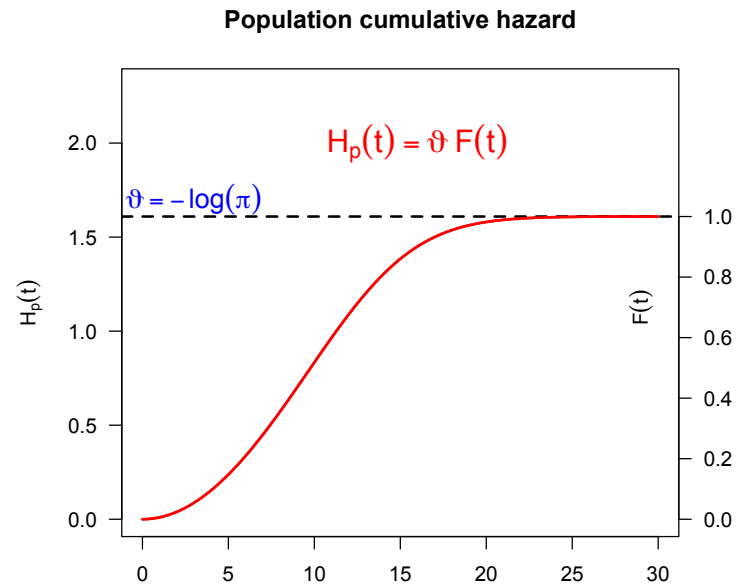
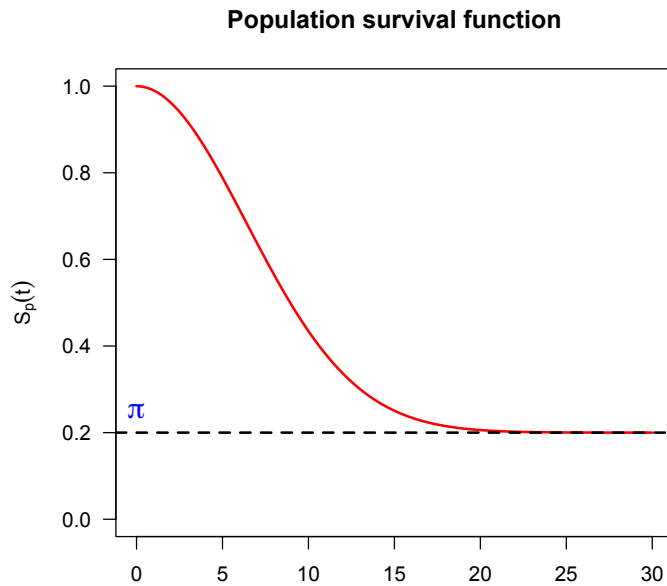


$$\text{where } S_p(t) = \exp(-H_p(t))$$

- The existence of a cured fraction  $\pi$  implies that the *cumulative hazard*  $H_p(t)$  at the population level is **bounded**: it remains constant when  $t \geq T$  :

$$t \geq T \implies H_p(t) = -\log S_p(+\infty) = -\log \pi = \vartheta$$

$$H_p(t) = \vartheta \times F(t)$$



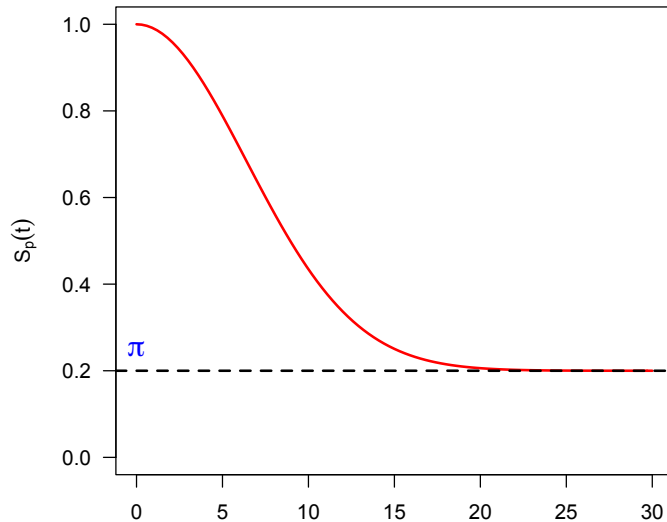
- Given that  $H_p(t)$  is a non decreasing function with  $H_p(0) = 0$ , one has

$$H_p(t) = \vartheta \times F(t)$$

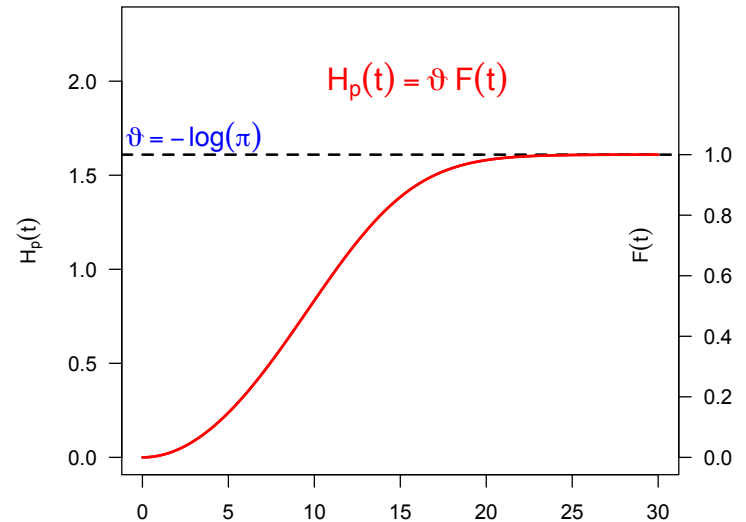
where  $F(\cdot)$  is a c.d.f. such that  $F(0) = 0$  and  $F(T) = 1$ .

$$S_p(t) = \exp \{ -\vartheta F(t) \}$$

Population survival function



Population cumulative hazard



- Cure fraction  $\implies \boxed{S_p(t) = \exp \{ -\vartheta F(t) \}}$  with  $P(\text{cured}) = \pi = \exp(-\vartheta)$
- Note:  $F(t)$  is the *normalized cumulative hazard*. It *governs* the *dynamics* of  $H_p(t)$  (i.e. *short-term survival*) towards  $\vartheta$ .
- The *bounded cumulative hazard model* or *promotion time cure model*, was initially motivated by a stochastic model of tumor latency (Yakovlev & Tsodikov 1996).

# Covariate inclusion in the promotion time model

- $H_p(t) = \vartheta(\mathbf{v}) \times F(t|\tilde{\mathbf{v}})$ 
  - ▷  $\mathbf{v} = (\mathbf{z}, \mathbf{x})$ : covariates influencing the cure probability (long-term survival);
  - ▷  $\tilde{\mathbf{v}} = (\tilde{\mathbf{z}}, \tilde{\mathbf{x}})$ : covariates influencing the hazard dynamics (short-term survival).
- One or more covariates could be shared by  $\mathbf{v}$  and  $\tilde{\mathbf{v}}$ .

- Where is  $H_p(t)$  going ?

$$\vartheta(\mathbf{v}) = -\log \underbrace{\pi(\mathbf{v})}_{\text{P(cure)}} = \exp\{\eta_{\vartheta}(\mathbf{v})\} \quad \text{where} \quad \eta_{\vartheta}(\mathbf{v}) = \beta_0 + \boldsymbol{\beta}^{\top} \mathbf{z} + \sum_j \underbrace{f_j(x_j)}_{\text{P-splines}}$$

- How fast does  $H_p(t)$  go to  $\vartheta$  ?

$$F(t|\tilde{\mathbf{v}}) = 1 - \underbrace{S_0(t)}_{\text{P-splines}} \exp(\eta_F(\tilde{\mathbf{v}})) \quad \text{where} \quad \eta_F(\tilde{\mathbf{v}}) = \boldsymbol{\gamma}^{\top} \tilde{\mathbf{z}} + \sum_j \underbrace{\tilde{f}_j(\tilde{x}_j)}_{\text{P-splines}}$$

- Then, the population survival function becomes

$$S_p(t|\mathbf{v}, \tilde{\mathbf{v}}) = \exp\{-H_p(t)\} = \exp\{-\vartheta(\mathbf{v})F(t|\tilde{\mathbf{v}})\}$$

# Data from German pension registers

- **Event** of interest: **age at 1st pregnancy** (1st birth minus 9 months).
  - Random sample of **15 248 women** from 'West' Germany.
  - **5 cohorts**: 1950-54, 1955-59, 1960-64, 1965-69 and 1970-74.
  - **Monthly** information from **Age 20 till 45** (maximum) on:
    - ▷ Occupation status ('Educ' ; 'Unemployed' ; 'Employed' ; 'Other') ;
    - ▷ Relative earnings (with 1.00 = Average gross earnings in a given year).
- ⇒ **time-varying** covariates !!
- Possibility of right censoring (i.e. no baby after a follow-up ending before age 45):

Cohort	<i>n</i>	Person months	Mother by Age 45		
			Yes	No	Right-cens.
1950-54	2423	277668	1900 (78.4%)	519 (21.4%)	4 (0.2%)
1955-59	2388	313692	1762 (73.8%)	613 (25.7%)	13 (0.5%)
1960-64	3029	431144	2186 (72.2%)	797 (26.3%)	46 (1.5%)
1965-69	3385	511119	2410 (71.2%)	887 (26.2%)	88 (2.6%)
1970-74	4023	655489	2761 (68.6%)	1130 (28.1%)	132 (3.3%)



# Extension to time-varying covariates

- Remember that

$$\begin{aligned} S_p(t|\mathbf{v}, \tilde{\mathbf{v}}) &= \exp\{-H_p(t)\} = \exp\{-\vartheta(\mathbf{v})F(t|\tilde{\mathbf{v}})\} \\ \implies h_p(t|\mathbf{v}, \tilde{\mathbf{v}}) &= e^{\eta_{\vartheta}(\mathbf{v}) + \eta_F(\tilde{\mathbf{v}})} h_0(t) S_0(t)^{\exp(\eta_F(\tilde{\mathbf{v}}))} \end{aligned}$$

- The key idea is to start from the population hazard specification:

$$\begin{aligned} \text{Substitute: } \mathbf{v} &\longrightarrow \mathbf{v}(t) \quad ; \quad \tilde{\mathbf{v}} \longrightarrow \tilde{\mathbf{v}}(t) \\ \implies h_p(t|\mathbf{v}(t), \tilde{\mathbf{v}}(t)) &= e^{\eta_{\vartheta}(\mathbf{v}(t)) + \eta_F(\tilde{\mathbf{v}}(t))} h_0(t) S_0(t)^{\exp(\eta_F(\tilde{\mathbf{v}}(t)))} \end{aligned}$$

- Likelihood contribution** for woman  $i$  during month  $m$  (with  $t = m\Delta$ ):
  - ▶ No event:  $\exp\{-h_p(t|\mathbf{v}_i(t), \tilde{\mathbf{v}}_i(t))\Delta\}$
  - ▶ Event:  $h_p(t|\mathbf{v}_i(t), \tilde{\mathbf{v}}_i(t)) \exp\{-h_p(t|\mathbf{v}_i(t), \tilde{\mathbf{v}}_i(t))\Delta\}$
- Smoothness priors** on the splines parameters in  $f_0(t)$ ,  $f_j(t)$  and  $\tilde{f}_j(t)$ :

$$p(\boldsymbol{\phi}|\omega) \propto \exp\left(-.5\boldsymbol{\phi}'(\omega\mathbf{P}_0)\boldsymbol{\phi}\right)$$

$$p(\boldsymbol{\theta}_j|\lambda_j) \propto \exp\left(-.5\boldsymbol{\theta}_j'(\lambda_j\mathbf{P})\boldsymbol{\theta}_j\right) \quad ; \quad p(\tilde{\boldsymbol{\theta}}_j|\tilde{\lambda}_j) \propto \exp\left(-.5\tilde{\boldsymbol{\theta}}_j'(\tilde{\lambda}_j\tilde{\mathbf{P}})\tilde{\boldsymbol{\theta}}_j\right).$$

# Inference

**Iterate** the following steps till convergence:  $\tau = (\omega, \lambda, \tilde{\lambda})$ ;  $\zeta = (\phi, \beta_0, \beta, \psi, \gamma, \tilde{\psi})$

1. Estimation of the **regression** and **spline parameters**  $\zeta$  for given penalty parameters:

▷ **Fast Maximization** of the **conditional posterior** using Newton-Raphson:

$$\log p(\zeta | \tau, \mathcal{D}) = \ell - \underbrace{\frac{\omega}{2} \phi^\top \mathbf{P}_0 \phi}_{h_0(t)} - \underbrace{\frac{1}{2} (\psi - \mathbf{b})^\top \mathbf{K}_\lambda (\psi - \mathbf{b})}_{\beta_0, \beta, f_j(t) \ (j=1, \dots, J)} - \underbrace{\frac{1}{2} (\tilde{\psi} - \mathbf{g})^\top \tilde{\mathbf{K}}_{\tilde{\lambda}} (\tilde{\psi} - \mathbf{g})}_{\gamma, \tilde{f}_j(t) \ (j=1, \dots, \tilde{J})}$$

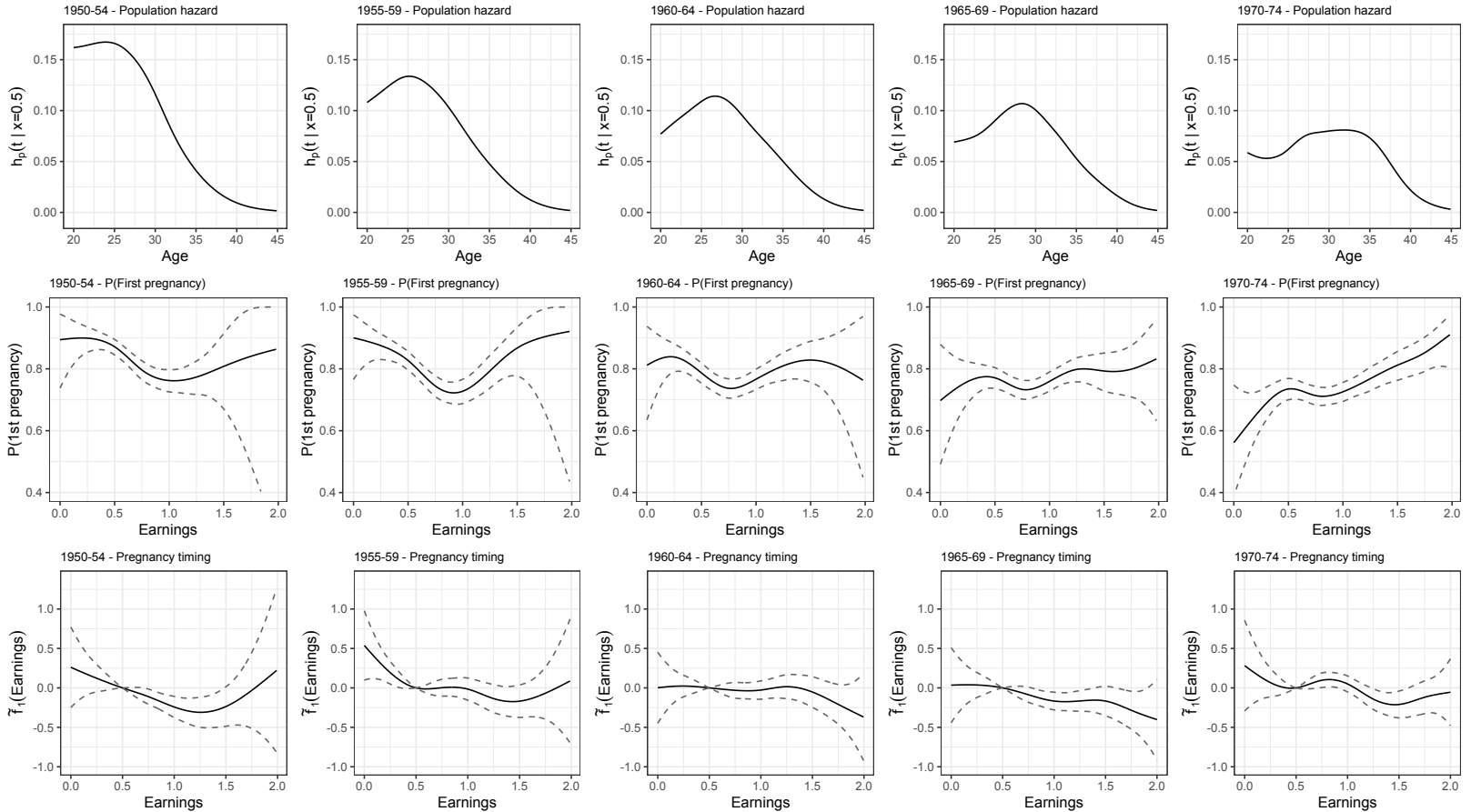
▷ **Explicit** forms exist for **gradients** and **Hessian** matrices.

2. **Selection** of the penalty parameters  $\tau = (\omega, \lambda, \tilde{\lambda})$ :

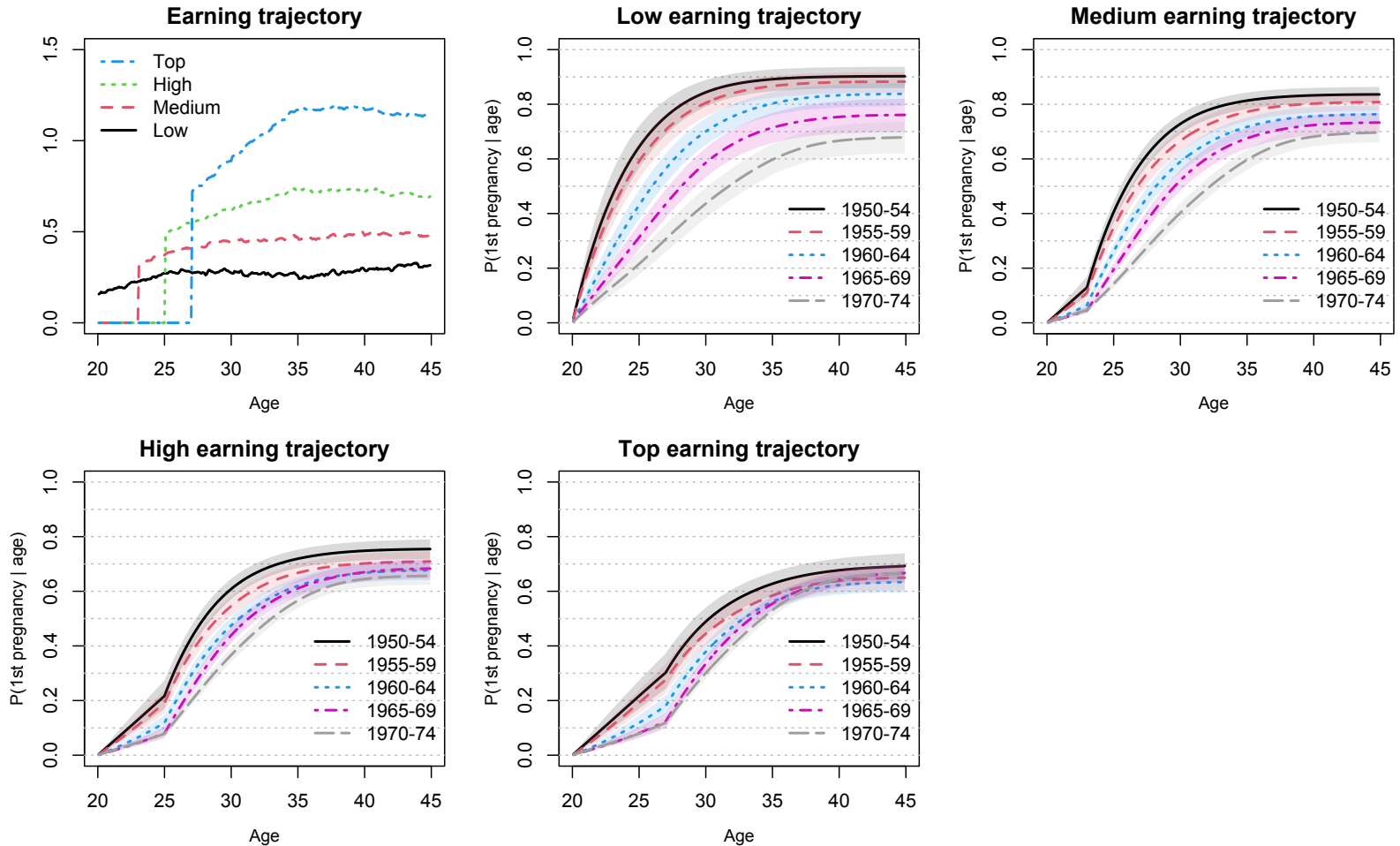
$$\begin{aligned} \text{Maximize } p(\tau | \mathcal{D}) &= \frac{p(\zeta, \tau | \mathcal{D})}{p(\zeta | \tau, \mathcal{D})} && \longleftarrow \text{Identity} \\ &\approx \frac{p(\zeta, \tau | \mathcal{D})}{p_G(\zeta | \tau, \mathcal{D})} && \longleftarrow \text{Laplace approximation} \\ &\propto p(\hat{\zeta}_\tau, \tau | \mathcal{D}) |\Sigma_\tau^{-1}|^{-1/2} && \longleftarrow \dots \text{evaluated at the MAP} \end{aligned}$$

↪ Maximization can be made **using** the **fixed point** method.

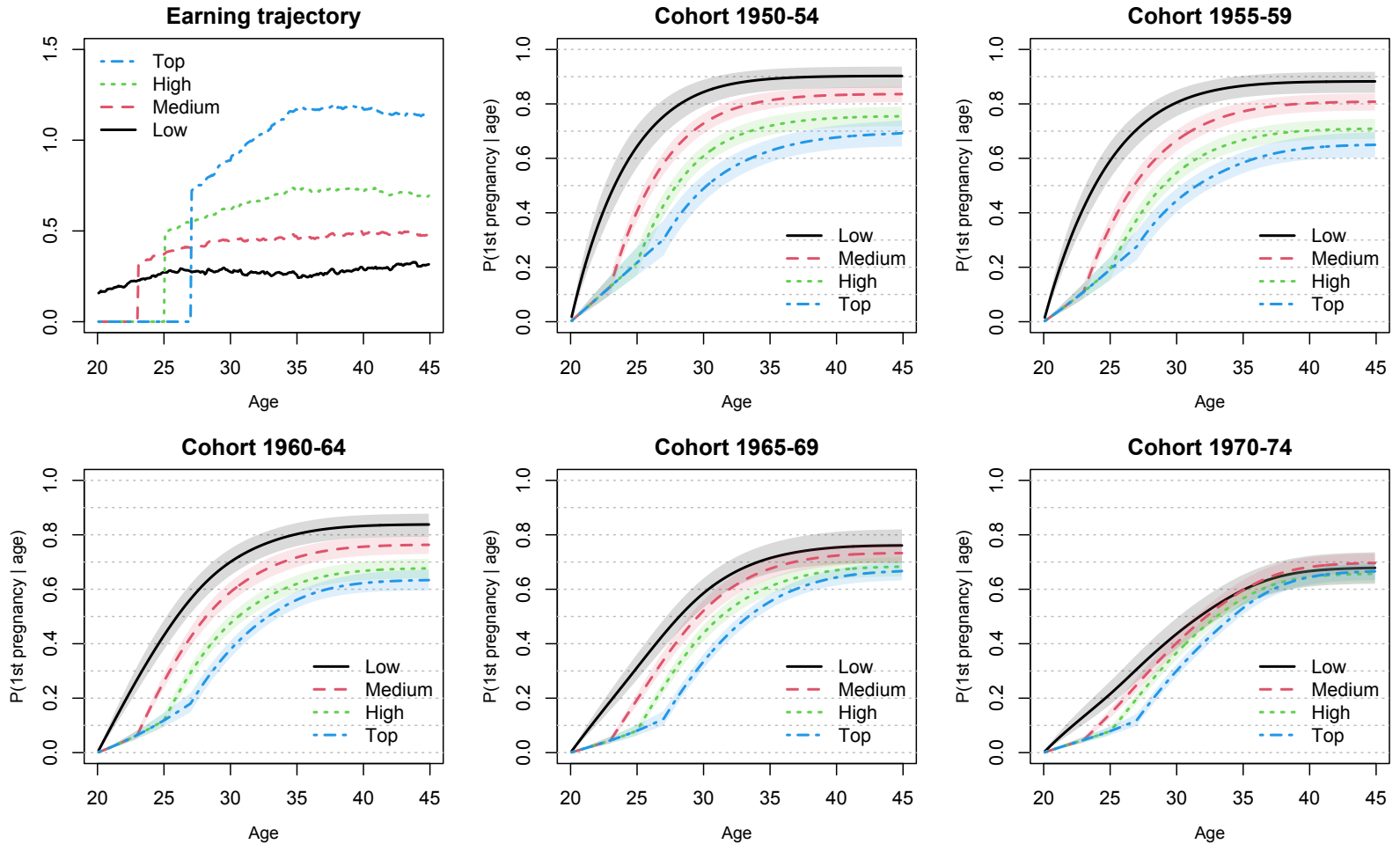
# Estimated model components per cohort



# Probability of pregnancy for prototypical earnings profiles



# ...from a cohort viewpoint



# Discussion

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- We described a very **flexible** cure survival model with:
  - ▷ Additive model for the **cure probability** ;
  - ▷ Flexible specification for the **event dynamics**.
- **Time-varying covariates** can be included in the two sub-models.
- Penalty parameters are selected to maximize their marginal posterior  
↔ **no over-fitting** (despite the rich information).
- Parameter interpretation in the presence of TV covariates is better understood by visualizing the estimated survival functions for given prototypical trajectories.  
Here, the employment status and earnings of a woman can vary over time in a non trivial way.
- R-package **tv cure** will soon be available on CRAN and on Github  
<https://github.com/plambertULiege/>

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